Small Area Estimation: Introduction to Survey Sampling

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Outline

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Survey Sampling

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Motivation

Motivation

- In the last set of slides, we motivated small area estimation, which is the estimation of some characteristic of a domain (area) based on sparse, or non-existent, data.
- In these slides we will discuss some common survey designs and design-based inference.
- This paradigm will be discussed for some simple designs.

Survey Sampling

Outline

Many national surveys employ stratified cluster sampling, also known as multistage sampling, so that's where we'd like to get to.

We will discuss:

- Simple Random Sampling (SRS)
- Stratified SRS
- Cluster sampling
- Multistage sampling

First, we briefly explain why taking account of the survey design (data collection process) is important.

Design-Based Inference

- We will focus on design-based inference: in this approach the population values of the variable of interest:

$$y_1, \ldots, y_N$$

  are viewed as fixed, while the indices of the individuals who are sampled are random.
- Imagine a population of size $N = 4$ and we sample $n = 2$
- There are 6 possible samples, with sampled unit indices in red and non-sampled in blue:

$$y_1, y_2, y_3, y_4$$
$$y_1, y_2, y_3, y_4$$
$$y_1, y_2, y_3, y_4$$
$$y_1, y_2, y_3, y_4$$
$$y_1, y_2, y_3, y_4$$
$$y_1, y_2, y_3, y_4$$

- Different designs are possible, and the probabilities we assign to each sample depend on which is used.

Design-Based Inference

Design-based inference is frequentist, so that properties are based on hypothetical replications of the data collection process; hence, we require a formal description of the replication process.

A complex random sample may be:

- Better than a simple random sample (SRS) in the sense of obtaining the same precision at lower cost, e.g., stratified sampling.
- May be worse in the sense of precision, but be required logistically, e.g., cluster sampling.

Probability Samples

Notation for random sampling, in a single population (and not distinguishing areas):

- N is population size.
- n is sample size.
- π_k is the sampling probability for a unit (which will often correspond to a person) k, $k = 1, \ldots, N$.

Random does not mean "equal chance", but means that the choice does not depend on variables/characteristics (either measured or unmeasured), except as explicitly stated via known sampling probabilities.

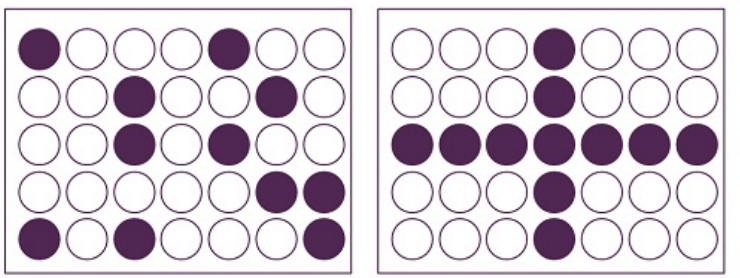
For example, in stratified random sampling, the probabilities of selection differ, in different strata.

Common sampling designs

- Simple random sampling: Select each individual with probability $\pi_k = n/N$.
- Stratified random sampling: Use information on each individual in the population to define strata h, and then sample n_h units independently within each stratum.
- Probability-proportional-to-size sampling: Given a variable related to the size of the sampling unit, Z_k, on each unit in the population, sample with probabilities $\pi_k \propto Z_k$.
- Cluster sampling: All units in the population are aggregated into larger units called clusters, known as primary sampling units (PSUs). Clusters are then sampled from this the set of PSUs, with units within these clusters being subsequently sampled.
- Multistage sampling: Stratified cluster sampling, with multiple levels of clustering.

Probability Samples

- The label probability sample is often used instead of random sample.

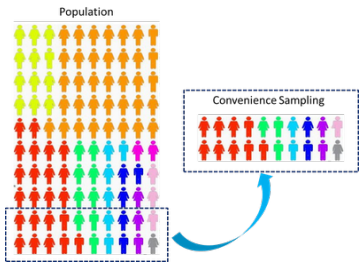


Probability Sampling Vs Non-Probability Sampling

- Non-probability samples cannot be analyzed with design-based approaches, because there are no π_k.

Non-probability sampling approaches include:

- Convenience sampling (e.g., asking for volunteers). Also known as accidental or haphazard sampling.



- Purposive (also known as judgmental) sampling in which a researcher uses their subject knowledge to select participants (e.g, selecting an "average" looking individual).
- Quota sampling in which quotas in different groups are satisfied (but unlike stratified sampling, probability sampling is not carried out, for example, the interviewer may choose friendly looking people!).

Probability Sampling and Inference

Probability Samples: Point Estimation

For design-based inference:

- To obtain an unbiased estimator, every individual k in the population needs to have a non-zero probability π_k of being sampled, $k = 1, \ldots, N$.
- To carry out inference, this probability π_k must be known only for every individual in the sample.
- So not needed for the unsampled individuals, which is key to implementation, since we will usually not know the sampling probabilities for those not sampled.

Probability Samples: Variance Estimation

For design-based inference:

- To obtain a form for the variance of an estimator: for every pair of units, k and l, in the sample, there must a non-zero probability of being sampled together, call this probability, π_{kl} for units k and l, $k = 1, \ldots, N$, $l = 1, \ldots, N$, $k \neq l$.
- The probability π_{kl} must be known for every pair in the sample.
- in practice, these are often approximated, or the variance is calculated via a resampling technique such as the jackknife.

Inference

- Suppose we are interested in a variable denoted y, with the population values being $y_1, \ldots, y_N$.
- Random variables will be represented by upper case letters, and constants by lower case letters.
- Finite population view: We have a population of size N and we are interested in characteristics of this population, for example, the mean:

$$\overline{y}_U = \frac{1}{N} \sum_{k=1}^{N} y_k.$$

Model-Based Inference

- Infinite population view: The population variables are drawn from a hypothetical distribution, with mean μ.
- In the model-based view, $Y_1, \ldots, Y_N$ are random variables and properties are defined with respect to $p(\cdot)$; often we say Y_k are independent and identically distributed (iid) from $p(\cdot)$.
- As an estimator of μ, we may take the sample mean:

$$\widehat{\mu} = \frac{1}{n} \sum_{k=1}^{n} Y_k.$$

- $\widehat{\mu}$ is a random variable because $Y_1, \ldots, Y_n$ are each random variables.
- Assume Y_k are iid observations from a distribution, $p(\cdot)$, with mean μ and variance σ^2.
- The sample mean is an unbiased estimator, and has variance σ^2/n.

- Unbiased estimator:

$$\begin{aligned} \mathsf{E}[\widehat{\mu}] &= \mathsf{E}\left[\frac{1}{n}\sum_{k=1}^{n} Y_k\right] = \frac{1}{n}\sum_{k=1}^{n} \underbrace{\mathsf{E}[Y_k]}_{=\mu} \\ &= \frac{1}{n}\sum_{k=1}^{n} \mu = \mu \end{aligned}$$

- Variance:

$$\begin{aligned} \mathsf{var}(\widehat{\mu}) &= \mathsf{var}\left(\frac{1}{n}\sum_{k=1}^{n} Y_k\right) \underbrace{=}_{\text{iid}} \frac{1}{n^2}\sum_{k=1}^{n} \underbrace{\mathsf{var}(Y_k)}_{=\sigma^2} \\ &= \frac{1}{n^2}\sum_{k=1}^{n} \sigma^2 = \frac{\sigma^2}{n} \end{aligned}$$

Model-Based Inference

- The variance σ^2 is unknown so we estimate by the unbiased estimator

$$s^2 = \frac{1}{n-1}\sum_{k=1}^{n}(y_k - \widehat{\mu})^2.$$

- A 95% asymptotic confidence interval is,

$$\widehat{\mu} \pm 1.96 \times \frac{s}{\sqrt{n}}.$$

- In practice, "asymptotic" means that n is sufficiently large that the sampling distribution of $\widehat{\mu}$ (i.e., it's distribution in hypothetical repeated samples) is close to normal.

Design-Based Inference

- In the design-based approach to inference the y values are treated as unknown but fixed.
- To emphasize: the y's are not viewed as random variables, so we write

  $$y_1, \ldots, y_N,$$

  and the randomness, with respect to which all procedures are assessed, is associated with the particular sample of individuals that is selected, call the random set of indices S.
- Minimal reliance on distributional assumptions.
- Sometimes referred to as inference under the randomization distribution.
- In general, the procedure for selecting the sample is under the control of the researcher.

Design-Based Inference

- Define design weights as

$$w_k = \frac{1}{\pi_k}.$$

- The basic estimator is the weighted mean (Horvitz and Thompson, 1952; Hájek, 1971)

$$\widehat{\overline{y}}_U = \frac{\sum_{k \in S} w_k y_k}{\sum_{k \in S} w_k}.$$

- This is an estimator of the finite population mean $\overline{y}_U$.
- So long as the weights are correctly calculated, and the sample size is not small, this estimator is appealing, though it may have high variance, if n is small.

The weighted mean is the basic **direct estimator** that is the first choice for SAE.

Complex Sampling Schemes

Simple Random Sample (SRS)

- The simplest probability sampling technique is simple random sampling without replacement.
- Suppose we wish to estimate the population mean in a particular population of size N.
- In everyday language: consider a population of size N; a random sample of size $n \leq N$ means that any subset of n people from the total number N is equally likely to be selected.

Simple Random Sample (SRS)

- We sample n people from N, choosing each person independently at random and with the same probability of being chosen:

$$\pi_k = \frac{n}{N},$$

  $k = 1, \ldots, N$.

- Since sampling without replacement the joint sampling probabilities are

$$\pi_{kl} = \frac{n}{N} \times \frac{n-1}{N-1}$$

  for $k, l = 1, \ldots, N$, $k \neq l$.

- In this situation:
  - The sample mean is an unbiased estimator.
  - The uncertainty, i.e. the variance, of the estimator can be easily estimated.
  - Unless n is quite close to N, the uncertainty does not depend on N, only on n.

The Indices are Random!

- **Example:** $N = 4, n = 2$ with SRS. There are 6 possibilities:

$$\{y_1, y_2\}, \quad \{y_1, y_3\}, \quad \{y_1, y_4\}, \quad \{y_2, y_3\}, \quad \{y_2, y_4\}, \quad \{y_3, y_4\}.$$

- The random variable describing this design is S, the set of indices of those selected.
- The sample space of S is

$$\{(1,2), \quad (1,3), \quad (1,4), \quad (2,3) \quad (2,4), \quad (3,4)\}$$

  and under SRS, the probability of sampling one of these possibilities is 1/6.
- The selection probabilities are

$$\pi_k = \Pr(\text{ individual } k \text{ in sample }) = \frac{3}{6} = \frac{1}{2}$$

  which is of course $\frac{n}{N}$.
- In general, we can work out the selection probabilities without enumerating all the possibilities!

Design-Based Inference

- Fundamental idea behind design-based inference: An individual with a sampling probability of π_k can be thought of as representing $w_k = 1/\pi_k$ individuals in the population.
- **Example:** in SRS each person selected represents $\frac{N}{n}$ people.
- The sum of the design weights,

$$\sum_{k \in S} w_k = n \times \frac{N}{n} = N,$$

  is the total population.
- Sometimes the population size may be unknown and the sum of the weights provides an unbiased estimator.
- In general, examination of the sum of the weights can be useful as if it far from the population size (if known) then it can be indicative of a problem with the calculation of the weights.

Estimator of $\bar{y}_U$ and Properties under SRS

- The weighted estimator is

$$\begin{aligned}\widehat{\bar{y}}_U &= \frac{\sum_{k\in S} w_k y_k}{\sum_{k\in S} w_k} \\ &= \frac{\sum_{k\in S} \frac{N}{n} y_k}{\sum_{k\in S} \frac{N}{n}} \\ &= \frac{\sum_{k\in S} y_k}{n} = \bar{y},\end{aligned}$$

  the sample mean, which is reassuring under SRS!

- This is an unbiased estimator, i.e.,

$$\mathrm{E}\left[\widehat{\bar{y}}_U\right] = \bar{y}_U,$$

  where we average over all possible samples we could have drawn, i.e., over S.

Unbiasedness

- For many designs: $\sum_{k \in S} w_k = N$ so we examine the estimator

$$\widehat{\overline{y}}_U = \frac{1}{N} \sum_{k \in S} w_k y_k.$$

- There's a neat trick in here, we introduce an indicator random variable of selection $I_k \sim \text{Bernoulli}(\pi_k)$:

$$\begin{aligned} \mathrm{E}\left[\widehat{\overline{y}}_U\right] &= \underbrace{\mathrm{E}\left[\frac{1}{N} \sum_{k \in S} w_k y_k\right]}_{S \text{ is random in here}} = \underbrace{\mathrm{E}\left[\frac{1}{N} \sum_{k=1}^{N} I_k w_k y_k\right]}_{I_k \text{ are random in here}} \\ &= \frac{1}{N} \sum_{k=1}^{N} \mathrm{E}\left[I_k\right] w_k y_k = \frac{1}{N} \sum_{i=1}^{N} \pi_k \frac{1}{\pi_k} y_k = \frac{1}{N} \sum_{i=1}^{N} y_k = \overline{y}_U \end{aligned}$$

Estimator of $\bar{y}_U$ and Properties under SRS

- It can be shown that the variance is

$$\text{var}(\bar{y}) = \left(1 - \frac{n}{N}\right) \frac{S^2}{n}, \tag{1}$$

where,

$$S^2 = \frac{1}{N-1} \sum_{k=1}^{N} (y_k - \bar{y}_U)^2.$$

- Contrast (1) with the model-based variance which is σ^2/n.
- The factor

$$1 - \frac{n}{N}$$

is the famous finite population correction (fpc) factor.

- Because we are estimating a finite population mean, the greater the sample size relative to the population size, the more information we have (relatively speaking), and so the smaller the variance.
- In the limit, if $n = N$ we have no uncertainty, because we know the population mean!

Estimator of $\bar{y}_U$ and Properties under SRS

- The variance of the estimator depends on the population variance S^2, is unknown, and we estimate using the unbiased estimator:

$$s^2 = \frac{1}{n-1} \sum_{k \in S} (y_k - \bar{y})^2.$$

- Substitution into (1) gives an unbiased estimator of the variance:

$$\widehat{\text{var}}(\bar{y}) = \left(1 - \frac{n}{N}\right) \frac{s^2}{n}. \tag{2}$$

- The standard error is

$$\text{SE}(\bar{y}) = \sqrt{\left(1 - \frac{n}{N}\right) \frac{s^2}{n}}.$$

- Note: S^2 is not a random variable but s^2 is.

Estimator of $\bar{y}_U$ and Properties under SRS

- If n, N and $N - n$ are "sufficiently large"[1], a 95% asymptotic confidence interval for $\bar{y}_U$ is

$$\bar{y} \pm 1.96 \times \sqrt{1 - \frac{n}{N}} \frac{s}{\sqrt{n}}. \tag{3}$$

- The interval given by (3) is random (across samples) because $\bar{y}$ and s^2 (the estimate of the variance) are random.
- In practice therefore, if $n \ll N$, we obtain the same confidence interval whether we take a design- or a model-based approach to inference (though the interpretation is different).

[1] so that the normal distribution provides a good approximation to the sampling distribution of the estimator

Two Inferential Settings

In a non-complex survey sampling setting:

- The data are (often implicitly) assumed to arise from simple random sample (SRS).
- The target of interest is a parameter in a hypothetical infinite population.

In a survey sampling setting:

- The data arise from a complex design, such as stratified two-stage unequal probability cluster sampling.
- The target of interest is a finite population characteristic.

In this second setting, statistical methods are required which acknowledge these two aspects.

Inference in SRS Setting

- When faced with 0/1 responses from a SRS, the usual approach is to assume an underlying risk parameter in area i is r_i.
- The risk r_i is the proportion of events of interest (e.g., neonatal deaths), in a hypothetical infinite "superpopulation" residing in area i.
- Let $Y_{ik} = 0/1$ be the indicator for the event k in area i with $k = 1, \ldots, n_i$, so that n_i is the number of samples in area i.
- In a model-based approach, it is often assumed that the total number of 1's, $Y_i = \sum_{k=1}^{n_i} Y_{ik}$, is distributed as Binomial(n_i, r_i).
- Under SRS in area i, the estimator is the simple mean

$$\widehat{r}_i^{\text{SRS}} = \frac{Y_i}{n_i} = \frac{1}{n_i} \sum_{k=1}^{n_i} Y_{ik}, \qquad \text{Std Err}(\widehat{r}_i^{\text{SRS}}) = \sqrt{\frac{\widehat{r}_i^{\text{SRS}}(1 - \widehat{r}_i^{\text{SRS}})}{n_i}}.$$

Stratified Sampling

Stratified Sampling

- Simple random samples are rarely taken in surveys because they are logistically difficult and there are more efficient designs for gaining the same precision at lower cost.
- Stratified random sampling is one way of increasing precision and involves dividing the population into groups called strata and drawing probability samples from within each one, with sampling from different strata being carried out independently.
- An important practical consideration of whether stratified sampling can be carried out is whether stratum membership is known for every individual in the population, i.e., we need a sampling frame containing the strata variable.

Rationale for Stratified Sampling

Lohr (2010, Section 3.1) provides a good discussion of the benefits of stratified sampling, we summarize here.

- Protection from the possibility of a "really bad sample", i.e., very few or zero samples in certain stratum giving an unrepresentative sample.
- Obtain known precision required for subgroups (domains) of the population – this is usual for the DHS.
- For example, from the Kenya DHS sampling manual (Kenya National Bureau of Statistics, 2015):

  "The 2014 KDHS was designed to produce representative estimates for most of the survey indicators at the national level, for urban and rural areas separately, at the regional (former provincial) level, and for selected indicators at the county level."

Rationale for Stratified Sampling

- Flexible since sampling frames can be constructed differently in different strata.
- For example, one may carry out different sampling in urban and rural areas.
- More precise estimates can be obtained if stratum can be found that are associated with the response of interest, for example, age and gender in studies of human disease.
- In a national study, the most natural form of sampling may be based on geographical regions.
- Due to the independent sampling in different stratum, variance estimation is straightforward, as long as within-stratum sampling variance estimators are available.

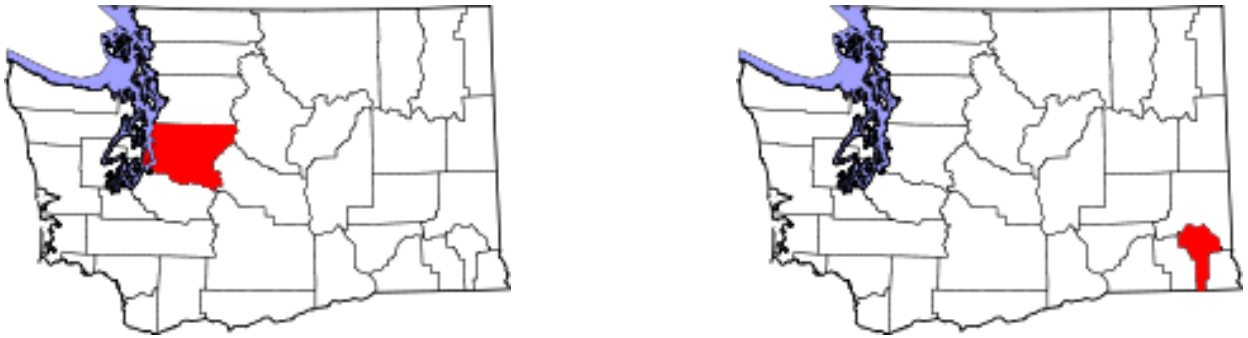
Example: Washington State

- According to the census there were 2,629,126 households in Washington State in the period 2009–2013.
- Consider a simple random sample (SRS) of 2000 households, so that each household has a

$$\frac{2000}{2629126} = 0.00076,$$

  chance of selection.
- Suppose we wish to estimate characteristics of household in all 39 counties of WA.

Example: Washington State



- King (highlighted left) and Garfield (highlighted right) counties had 802,606 and 970 households so that under SRS we will have, on average, about 610 households sampled from King County and about 0.74 from Garfield county.
- The probability of having no-one from Garfield County is about 22% (binomial experiment), and the probability of having more than one is about 45%.
- If we took exactly 610 from King and 1 (rounding up) from Garfield we have an example of proportional allocation, which would not be a good idea given the objective here.
- Stratified sampling would allow control of the number of samples in each county.

Notation

- Stratum levels are denoted $h = 1, \ldots, H$, so H in total.
- Let $N_1, \ldots, N_H$ be the known population totals in the stratum with

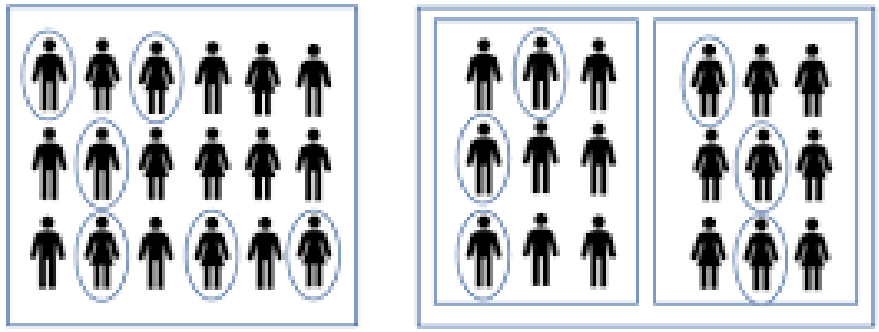
$$N_1 + \cdots + N_H = N,$$

so that N is the total size of the population.

- In stratified simple random sampling, the simplest from of stratified sampling, we take a SRS from each stratum with n_h samples being randomly taken from stratum h, so that the total sample size is

$$n_1 + \cdots + n_H = n.$$

Figure 1: Comparison of Simple Random Sampling to Stratified Random Sampling



Visual of Simple Random Sampling: Selection of 6 out of 18 People

Visual of Stratified Random Sampling: Selection of 3 out of 9 Men and 3 out of 9 Women

- We can view stratified SRS as carrying out SRS in each of the H stratum; we let S_h represent the probability sample in stratum h.
- We also let S refer to the overall probability sample.

Estimators

- The sampling probabilities for unit k in strata h are

$$\pi_{hk} = \frac{n_h}{N_h},$$

which do not depend on k.

- Therefore the design weights are

$$w_{hk} = \frac{N_h}{n_h}.$$

- Note that:

$$\sum_{h=1}^{H} \sum_{k \in S_h} w_{hk} = \sum_{h=1}^{H} \sum_{k \in S_h} \frac{N_h}{n_h} = \sum_{h=1}^{H} n_h \frac{N_h}{n_h} = N,$$

so that summing over the weights recovers the population size.

- Weighted estimator:

$$\widehat{\overline{y}}_U = \frac{\sum_{h=1}^{H}\sum_{k\in S_h} w_{hk}y_{hk}}{\sum_{h=1}^{H}\sum_{k\in S_h} w_{hk}} = \sum_{h=1}^{H}\frac{N_h}{N}\overline{y}_h$$

where

$$\overline{y}_h = \frac{\sum_{k\in S_h} y_{hk}}{n_h}.$$

- Since we are sampling independently from each stratum using SRS, we have[2]

$$\text{var}(\widehat{\overline{y}}_U) = \sum_{h=1}^{H}\left(\frac{N_h}{N}\right)^2\left(1-\frac{n_h}{N_h}\right)\frac{s_h^2}{n_h}, \tag{4}$$

where the within stratum variances are:

$$s_h^2 = \frac{1}{n_h-1}\sum_{k\in S_h}(y_{hk}-\overline{y}_h)^2.$$

[2] using the variance formula for SRS

Weighted Estimation

Recall: The weight w_k can be thought of as the number of people in the population represented by sampled person k.

Example 1: Simple Random Sampling

Suppose an area contains 1000 people:

- Using simple random sampling (SRS), 100 people are sampled.
- Sampled individuals have weight $w_k = 1/\pi_k = 1000/100 = 10$.

Example 2: Stratified Simple Random Sampling

Suppose an area contains 1000 people, 200 urban and 800 rural.

- Using stratified SRS, 50 urban and 50 rural individuals are sampled.
- Urban sampled individuals have weight $w_k = 1/\pi_k = 200/50 = 4$.
- Rural sampled individuals have weight $w_k = 1/\pi_k = 800/50 = 16$.

Example 2: Stratified Simple Random Sampling

Suppose an area contains 1000 people, 200 urban and 800 rural:

- Urban risk = 0.1.
- Rural risk = 0.2.
- **True risk = 0.18.**

Take a stratified SRS, 50 urban and 50 rural individuals sampled:

- Urban sampled individuals have weight 4; 5 cases out of 50.
- Rural sampled individuals have weight 16; 10 cases out of 50.
- **Simple mean is** $15/100 = 0.15 \neq 0.18$**.**
- **Weighted mean is**

$$\frac{4 \times 5 + 16 \times 10}{4 \times 50 + 16 \times 50} = \frac{180}{1000} = 0.18.$$

Cluster Sampling

Motivation for Cluster Sampling

For logistical reasons, cluster sampling is an extremely common design that is often used for government surveys.

Two main reasons for the use of cluster sampling:

- A sampling frame for the population of interest does not exist, i.e., no list of population units.
- The population units have a large geographical spread and so direct sampling is not logistically feasible to implement for in-person interviews.
- It is far more cost effective (in terms of travel costs, etc.) to cluster sample.

Terminology

- In single-stage cluster sampling or one-stage cluster sampling, the population is grouped into subpopulations (as with stratified sampling) and a probability sample of these clusters is taken, and every unit within the selected clusters is surveyed.
- In one-stage cluster sampling either all or none of the elements that compose a cluster (PSU) are in the sample.
- The subpopulations are known as clusters or primary sampling units (PSUs).
- In two-stage cluster sampling, rather than sample all units within a PSU, a further cluster sample is taken; the possible groups to select within clusters are known as secondary sampling units (SSUs).
- This can clearly be extended to multistage cluster sampling.

Differences Between Cluster and Stratified sampling

Stratified Random Sampling	One-Stage Cluster Sampling
A sample is taken from every stratum	Observe all elements only within the sampled clusters
Variance of estimate of $\overline{y}_U$ depends on within strata variability	The cluster is the sampling unit and the more clusters sampled the smaller the variance – which depends primarily on between cluster means
For greatest precision, we want low within-strata variability but large between-strata variability	For greatest precision, high within-cluster variability and similar cluster means.
Precision generally better than SRS	Precision generally worse than SRS



Stratified Sampling Vs Cluster Sampling

Heterogeneity

- The reason that cluster sampling loses efficiency over SRS is that within clusters we only gain partial information from additional sampling within the same cluster, since within clusters two individuals tend to be more similar than two individuals within different clusters.
- The similarity of elements within clusters is due to unobserved (or unmodeled) variables.
- The design effect (deff) is often to summarize the effect on the variance of the design:

$$\text{deff} = \frac{\text{Variance of estimator under design}}{\text{Variance of estimator under SRS}},$$

where in the denominator we use the same number of observations as in the complex design in the numerator.

Estimation for One-Stage Cluster Sampling

- We suppose that a SRS of n PSUs is taken.
- The probability of sampling a PSU is n/N, and since all the SSUs are sampled in each selected PSU we have selection probabilities and design weights:

$$\pi_{ik} = \Pr(\text{ SSU } k \text{ in cluster } i \text{ is selected }) = \frac{n}{N}$$

$$w_{ik} = \text{Design weight for SSU } k \text{ in cluster } i = \frac{N}{n}.$$

Let S represent the set of sampled clusters.

Estimation for One-Stage Cluster Sampling

- Let $M_0 = \sum_{i=1}^{N} M_i$ be the total number of secondary sampling units (SSUs), i.e., elements in the population, so the population mean is

$$\bar{y}_U = \frac{1}{M_0} \sum_{i=1}^{N} \sum_{k=1}^{M_i} y_{ik}$$

- An unbiased estimator is

$$\widehat{\bar{y}}_U = \frac{\sum_{i \in s} \sum_{k \in s_i} w_{ik} y_{ik}}{M_0}.$$

- Then,

$$\widehat{\text{var}}(\widehat{\bar{y}}_U) = \frac{N^2}{M_0^2} \left(1 - \frac{n}{N}\right) \frac{s_T^2}{n}$$

where s_T^2 is the estimated variance of the PSU totals.

Two-Stage Cluster Sampling with Equal-Probability Sampling

It may be wasteful to measure all SSUs in the selected PSUs, since the units may be very similar and so there are diminishing returns on the amount of information we obtain.

We discuss the equal-probability two stage cluster design:

1. Select a SRS of n PSUs from the population of N PSUs.
2. Select a SRS of m_i SSUs from each selected PSU, the probability sample collected will be denoted S_i.

Two-Stage Cluster Sampling Weights

- The selection probabilities are:

$$\begin{aligned}\Pr(\ k\text{-th SSU in } i\text{-th PSU selected }) &= \Pr(\ i\text{-th PSU selected }) \\ &\times \Pr(\ k\text{-th SSU} \mid i\text{-th PSU selected }) \\ &= \frac{n}{N} \times \frac{m_i}{M_i}\end{aligned}$$

- Hence, the weights are

$$w_{ik} = \pi_{ik}^{-1} = \frac{N}{n} \times \frac{M_i}{m_i}.$$

- An unbiased estimator is

$$\widehat{\bar{y}}_U = \frac{\sum_{i \in s} \sum_{k \in s_i} w_{ik} y_{ik}}{M_0}.$$

- Variance calculation is not trivial, and requires more than knowledge of the weights.

Variance Estimation for Two-Stage Cluster Sampling

- In contrast to one-stage cluster sampling we have to acknowledge the uncertainty in both stages of sampling; in one-stage cluster sampling the totals t_i are known in the sampled PSUs, whereas in two stage sampling we have estimates $\widehat{t}_i$.
- In Lohr (2010, Chapter 6) it is shown that

$$\text{var}(\widehat{\bar{y}}_U) = \frac{1}{M_0^2}\left[\underbrace{N^2\left(1-\frac{n}{N}\right)\frac{s_T^2}{n}}_{\text{One-stage cluster variance}} + \underbrace{\frac{N}{n}\sum_{i\in S}\left(1-\frac{m_i}{M_i}\right)M_i^2\frac{s_i^2}{m_i}}_{\text{Two-stage cluster variance}}\right] \tag{5}$$

  where
  - s_T^2 is the estimated variance of the cluster totals,
  - s_i^2 is the estimated variance within the i-th PSU.
- In most software packages, the second term in (5) is ignored, since it is small when compared to the first term, when N is large.

The Jackknife

- The jackknife is a very general technique for calculating the variance of an estimator.
- The basic idea is to delete portions of the data, and then fit the model on the remainder – if one repeats this process for different portions, one can empirically obtain the distribution of the estimator.
- The key is to select the carefully select the portion of the data so that the design is respected.
- We describe in the context of multistage cluster sampling.
- Observations within a PSU should be kept together when constructing the data portions, which preserves the dependence among observations in the same PSU.

The Jackknife for Multistage Cluster Sampling

- Assume we have H strata and n_h PSUs in strata h, and assume PSUs are chosen with replacement.
- To apply the jackknife, delete one PSU at a time.
- Let $\widehat{\mu}_{(hi)}$ be the estimator when PSU i of stratum h is omitted.
- To calculate $\widehat{\mu}_{(hi)}$ we define a new weight variable:

$$w_{k(hi)} = \begin{cases} w_{k(hi)} & \text{if observation } k \text{ is not in stratum } h \\ 0 & \text{if observation } k \text{ is in PSU } i \text{ of stratum } h \\ \frac{n_h}{n_h - 1} w_k & \text{if observation } k \text{ is not in PSU } i \text{ but in stratum } h \end{cases}$$

Then we can use the weights $w_{k(hi)}$ to calculate $\widehat{\mu}_{(hi)}$ and

$$\widehat{V}_{\text{JK}}(\widehat{\mu}) = \sum_{h=1}^{H} \frac{n_h - 1}{n_h} \sum_{i=1}^{n_h} (\widehat{\mu}_{(hi)} - \widehat{\mu})^2.$$

Multistage Sampling in the DHS

- A common design in national surveys is multistage sampling, in which cluster sampling is carried out within strata.
- DHS Program: Typically, 2-stage stratified cluster sampling:
  - Strata are urban/rural and region.
  - Enumeration Areas (EAs) sampled within strata (PSUs).
  - Households within EAs (SSUs).
- Weighted estimators are used and a common approach to variance estimation is the jackknife (Pedersen and Liu, 2012)
- In later lectures, we will show how model-based inference can be carried out for the DHS.

Acknowledging the Design: Stratification

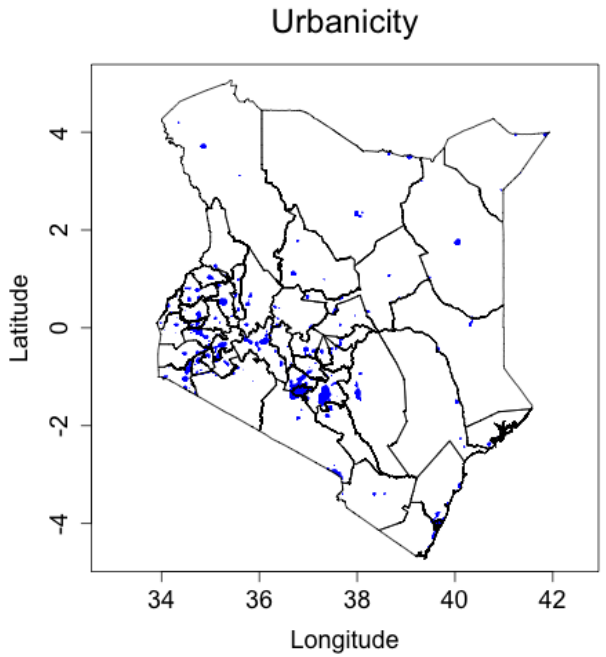


Figure 1: In the DHS, stratification is based on counties (the solid lines) and on a binary urban/rural variable (urban indicated in blue, the white is rural).

- Suppose we are interested in the proportion of women aged 20–29 who complete secondary education – this is much higher in urban areas
- If we oversample urban areas but ignore this when we analyze the data we will overestimate the fraction of women who complete secondary education, i.e., we will introduce bias.
- Taking into account of the stratification also reduces the variance of the estimator.
- In the design-based approach to inference, the stratification is accounted for via design weights.
- In the model-based approach to inference, the stratification is accounted for in the mean model.

Acknowledging the Design: Cluster Sampling

- The DHS also employs cluster sampling, in which multiple units (individuals) within the same cluster are interviewed.
- Units within the same cluster tend to be more similar than units in different clusters, which reduces the information content of the clustered sample, relative to independently sampled units.
- The dependence can be measured via the intraclass correlation coefficient.
- In the design-based approach to inference, the clustering is accounted for in the variance calculation that is carried out.
- In the model-based approach to inference, the clustering is accounted for by including a cluster-specific random effect in the model.

Modes of Inference

- Surveys can be analyzed using **design- and model-based inference**. In this lecture, the former will be focused upon.
- The target of inference are the set of means for areas indexed by i (e.g., Admin2 regions).
- Let y_{ik} be the binary indicator on the k-th unit sampled in area i, for $k \in S_i$ (the set of selected individuals) and $i = 1, \ldots, n$.

Design-Based Inference

- Labels S_i of sampled units are random.
- Responses y_{ik} are fixed.
- Asymptotic inference, perhaps using resampling.

Model-Based Inference

- Condition on units that are actually sampled.
- Responses Y_{ik} are random.
- Exact inference, conditional on model.

Model-Based Inference

Suppose we carry out stratified cluster sampling, with one-stage of clusters, and the outcome is continuous.

Let y_{ck} be the outcome from sampling unit k in sampled cluster c, and $\boldsymbol{s}_c$ the location of cluster c,

Suppose the data were collected within two strata, urban and rural.

A model-based approach to inference might begin with

$$Y_{ck} = \alpha + \gamma I(\boldsymbol{s}_c \in \text{ rural }) + \epsilon_c + \upsilon_{ck},$$

where

- α is the mean for urban and $\alpha + \gamma$ is the mean for rural.
- within-cluster dependence is modeled via the random effect $\epsilon_c \sim_{iid} \text{N}(0, \sigma^2_\epsilon)$.
- Measurement error is $\upsilon_{ck} \sim_{iid} \text{N}(0, \sigma^2_\upsilon)$.

Discussion

Discussion

- Understanding the motivation for and differences between stratified sampling and cluster sampling is important.
- Understanding the derivation of the variance formulas is less important.
- **Important to recognize that we can't ignore the sampling design.**

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